

# Fixed Income Analysis

## Calibration in lattice models — Part II

Calibration to the initial volatility structure

Pitfalls in volatility calibrations

Mean-reverting log-normal models (Black-Karasinski)

Brownian-path independent (BPI) models

Trinomial lattice models (Hull-White)

Jesper Lund

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### The volatility structure

- **Definition:** volatility of zero-coupon yields as a function of time to maturity.
- In a one-factor model, the zero-coupon rate is governed by

$$dR(t, T) = \mu_R(t, T)dt + \sigma_R(t, T)R(t, T)dW_t. \quad (1)$$

- Here,  $\sigma_R(t, T)$  is the **proportional** volatility, and  $\sigma_R(t, T)R(t, T)$  is the **basis-point** volatility.
- From Ito's lemma, the volatility structure is given by

$$\sigma_R(t, T)R(t, T) = \sigma(r_t) \frac{\partial R(t, T)}{\partial r} = \frac{-\sigma(r_t)}{(T-t)P(t, T)} \frac{\partial P(t, T)}{\partial r} \quad (2)$$

- Thus, the volatility structure depends on the effect of the short rate,  $r_t$ , on bond prices,  $P(t, T)$ .
- Mainly determined by the speed of **mean reversion**.

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## Calibration in the BDT model

- The BDT model is an approximation to the SDE

$$d \log r_t = \left\{ b(t) + \frac{\sigma'(t)}{\sigma(t)} \log r_t \right\} dt + \sigma(t) dW_t^Q. \quad (3)$$

- Note that the mean reversion coefficient is tied to  $\sigma(t)$ , the short-rate volatility at time  $t$ .
- Calibrating the BDT model to the initial yield **and** volatility curve:
  - The basic geometry of the the BDT binomial tree is unchanged.
  - The pair  $\{b(n\Delta), \sigma(n\Delta)\}$  is chosen to match the yield and volatility of the  $(n + 1)$ -period bond.
  - Alternative parameterization:  $\{r(n, 0), \delta_n\}$ , where  $r(n, 0)$  is the bottom node and  $\log \delta_n$  is the spacing for  $\log r(n, s)$ .
  - We have two equations in two unknowns, but they are easy to solve numerically (Newton-Raphson) when the forward-induction technique is used.

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## Pitfalls in volatility calibration – 1

- The **basic** problem can be illustrated for the extended Vasicek model (in continuous time)

$$dr_t = \kappa(t) \{ \theta(t) - r_t \} dt + \sigma(t) dW_t^Q. \quad (4)$$

- The pair  $\{ \theta(t), \kappa(t) \}$  is chosen to fit the initial (time  $t = 0$ ) yield and volatility curves.
- The initial forward-rate volatility structure is given by:

$$\sigma(0, T) = \sigma(0) e^{-\int_0^T \kappa(u) du}. \quad (5)$$

- Basis-point volatility structure for zero-coupon rates:

$$\sigma_R(0, T) = \frac{1}{T} \int_0^T \sigma(0, s) ds = \sigma(0) \int_0^T e^{-\int_0^s \kappa(u) du} ds. \quad (6)$$

- Note that the shape only depends on  $\kappa(t)$ , for  $0 \leq t \leq T$ .

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## Pitfalls in volatility calibration – 2

- At time  $t$ , we have a new volatility structure, and the payoffs from derivatives (e.g., call options) depend on the time  $t$  volatilities.
- **Caveat:** in a Markovian model, the new volatility structure is completely determined from the initial volatility structure.
- New forward-rate volatilities:

$$\sigma(t, T) = \sigma(t) e^{-\int_t^T \kappa(u) du} = \sigma(t) \frac{\sigma(0, T)}{\sigma(0, t)}. \quad (7)$$

- New volatility structure for zero-coupon rates:

$$\sigma_R(t, T) = \frac{\sigma(t)}{T-t} \cdot \frac{T\sigma_R(0, T) - t\sigma_R(0, t)}{\sigma(0, t)} \quad (8)$$

- Apart from  $\sigma(t)$ , which is common for all maturities, (7) and (8) only depend on the initial volatility structure.

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## Pitfalls in volatility calibration – 3

- Constraining the evolution of the volatility structure in this way **could** have undesirable effects on derivatives prices.
- If the current volatility curve is humped, the future curve will be steeply downward sloping, although flat eventually.
- Hull and White's recommendation: do **not** calibrate the model to the volatility structure (if anything, use cap prices instead).
- Additional problems for the BDT model:
  - Minor problem: no analytical solution for bond prices, so the dependencies are more difficult to analyze (and understand).
  - Major problem: mean reversion is tied to the future short-rate volatilities since  $\kappa(t) = -\sigma'(t)/\sigma(t)$  in the BDT model.
- If the volatility structure is downward sloping (normal situation), we need BDT's  $\sigma(t)$  to be decreasing in  $t$  — and this is an **unrealistic** property (in general).

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## Pitfalls in volatility calibration – 4

- Illustration of the BDT volatility problem, taken from a paper by Simon Schultz and Per Sørensen, “Pricing Caps and Floors,” *Finans/Invest*, 4/93 [in Danish].
- Market prices (bid-ask) and three model prices for interest-rate caps (all prices are relative to bid-ask midpoint).
- The BDT and Hull-White (extended) Vasicek models are calibrated to a downward-sloping volatility structure.
- Since  $\sigma(t)$  in the BDT model is decreasing over time, the BDT cap prices generally are **too low** in the table below.

Maturity	Bid	Ask	BDT	Vasicek	Black-76
2	0.95	1.05	0.94	1.00	0.97
3	0.97	1.03	0.90	1.00	0.98
4	0.97	1.03	0.92	1.00	0.98
5	0.97	1.03	0.94	0.99	0.97

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## Mean reversion in log-normal models

- Black and Karasinski (1991) relax the BDT restriction on the mean reversion coefficient,

$$d \log r_t = \{b(t) - \phi(t) \log r_t\} dt + \sigma(t) dW_t^Q. \quad (9)$$

- The model (9) **cannot** be implemented in a recombining binomial tree with constant time steps and probabilities  $\theta(n, s) = 0.5$ .
- There are three possible modifications of the tree which allow for (arbitrary) mean reversion:
  1. Non-constant time steps,  $\Delta_n$ , with  $\theta(n, s) = 0.5$ . This is suggested by Black and Karasinski (1991). The main disadvantage is that the **spacing declines** over time, and we generally want the opposite (if anything).
  2. Non-constant probabilities of an up-move, but with constant time steps. We match the expected change in  $r$  (mean reversion) at node  $(n, s)$  by adjusting  $\theta(n, s)$ . The disadvantage is **slower convergence**.
  3. Trinomial trees (Hull-White) which have three branches at each node.

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## Brownian-path independence (BPI)

- Why is it possible to construct a **simple** binomial tree — with constant time steps and  $\theta(n, s) = 0.5$  — in the BDT case?
- The solution to the BDT SDE (3) can be written as

$$\begin{aligned}\log r_t &= \frac{\sigma(t)}{\sigma(0)} \log r_0 + \sigma(t) \int_0^t \frac{b(s)}{\sigma(s)} ds + \sigma(t) W_t^Q \\ &\equiv B(t) + \sigma(t) W_t^Q\end{aligned}\quad (10)$$

- The BDT model modifies the Brownian motion tree only by scaling [through  $\log(\delta_n) = c \cdot \sigma(t)$ ] and the bottom node,  $r(n, 0)$ .
- In most **BPI models**, the short rate has the general form:

$$r_t = F\left(B(t) + \sigma(t) W_t^Q\right), \quad (11)$$

for some function  $F(x)$ . Note that BDT has  $F(x) = \exp(x)$ .

- Simple binomial trees requires a BPI model [Jamshidian (1991)].

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## Trinomial lattices – 1

- Introduced by Hull and White (1993, 1994).
- Extended Vasicek model with  $\kappa(t) = \kappa$  and  $\sigma(t) = \sigma$ ,

$$dr_t = \kappa \{\theta(t) - r_t\} dt + \sigma dW_t^Q. \quad (12)$$

- We rewrite (12) as

$$r_t = \alpha(t) + x_t \quad (13)$$

$$\alpha(t) = e^{-\kappa t} r_0 + \int_0^t e^{-\kappa(t-s)} \kappa \theta(s) ds \quad (14)$$

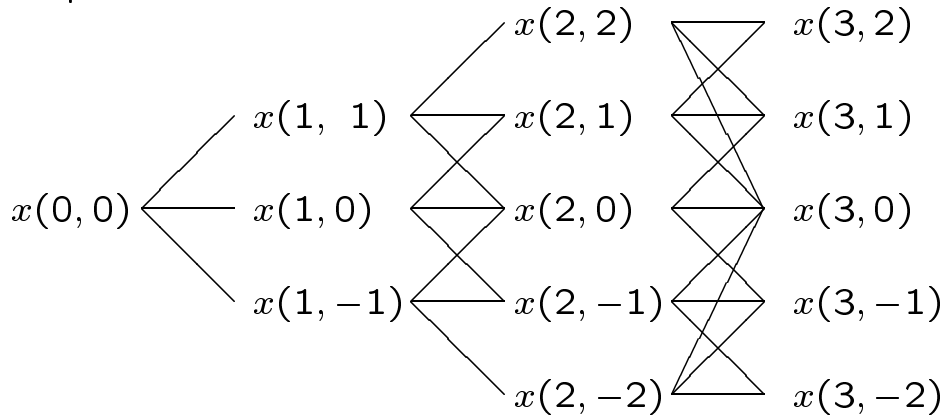
$$dx_t = -\kappa x_t dt + \sigma dW_t^Q, \quad \text{with } x_0 = 0. \quad (15)$$

- First step: build a trinomial tree for  $x_t$ . This tree is symmetric around  $x = 0$ , and the geometry depends only on  $\kappa$  and  $\sigma$ .
- Second step: calibrate the time-dependent parameters,  $\alpha_i = \alpha(t_i)$ , to match the initial term structure.

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## Trinomial lattices – 2

- Three-period trinomial tree:



- **Comment 1:** the numbering scheme,  $x(i, j)$ , is different from binomial case. The center node has  $j = 0$  for all times  $i$ .
- **Comment 2:** there is a different branching scheme for low and high values of  $j$  (node number). The purpose is accommodating mean reversion, while retaining positive probabilities.

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## Trinomial lattices – 3

- First and second moments of the SDE for  $x_t$ :

$$E_t [x_{t+\Delta} - x_t] \equiv Mx_t = (e^{-\kappa\Delta} - 1)x_t \approx -\kappa\Delta x_t \quad (16)$$

$$\text{Var}_t [x_{t+\Delta} - x_t] \equiv V = \frac{\sigma^2}{2\kappa} (1 - e^{-2\kappa\Delta}) \approx \sigma^2\Delta \quad (17)$$

- The two approximations follow from  $\exp(z) \approx 1 + z$ .
- The trinomial model has a constant time step, denoted  $\Delta$ .
- The spacing on the  $x$ -axis is specified as  $\Delta x = \sqrt{3V}$ .
- This means that  $x(i, j) = j\Delta x$ , for  $-n_i \leq j \leq n_i$ .
- From each node,  $(i, j)$ , there are branches to three nodes with probabilities:  $p_u$  (top node),  $p_m$  (mid mode), and  $p_d$  (low node).

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## Trinomial lattices – 4

- We look at a “normal” node  $(i, j)$  — not special branching.
- The probabilities  $p_u$ ,  $p_m$  and  $p_d$  are chosen in order to satisfy

$$p_u \Delta x - p_d \Delta = Mj \Delta x \quad (18)$$

$$p_u (\Delta x)^2 + p_d (\Delta x)^2 = V + M^2 j^2 (\Delta x)^2 \quad (19)$$

$$p_u + p_m + p_d = 1 \quad (20)$$

- That is, we match the moments of  $(x_{t+\Delta} - x_t)$ , see (16) and (17).
- Since  $V = (\Delta x)^2/3$ , the solution is easily found as

$$p_u = \frac{1}{6} + \frac{j^2 M^2 + jM}{2} \quad (21)$$

$$p_m = \frac{2}{3} - j^2 M^2 \quad (22)$$

$$p_d = \frac{1}{6} + \frac{j^2 M^2 - jM}{2} \quad (23)$$

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## Trinomial lattices – 5

- Note that the probabilities are independent of the initial term structure. Apart from  $j$ , they only depend on  $\kappa$  and  $\Delta$ .
- When branching out from the special top and bottom nodes, similar formulas apply — see Hull (1997) or Hull & White (1994).
- This completes the first step, setting up the nodes of the trinomial lattice.
- Second step: let  $r(i, j) = \alpha_i + x(i, j)$ , and **calibrate**  $\alpha_i$  recursively so that the  $(i + 1)$ -period bond price is matched exactly.
- As in the BDT model, the calibration is done with forward induction and Arrow-Debreu prices.
- We start with  $\alpha_0 = r(0, 0) = -\log P(1)/\Delta$ .
- **Define**  $p(i, j) = \exp[-r(i, j)\Delta] = \exp[-\alpha_i \Delta] \cdot \exp[-j(\Delta x)\Delta]$ .

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## Trinomial lattices – 6

- Assume that we have computed  $\alpha_{m-1}$ , where  $m \geq 1$ .
- First, we use the forward equation to compute  $G(m, j)$ ,

$$G(m, j) = \sum_{k=-(n_{m-1})}^{n_{m-1}} q(k, j)p(m-1, k)G(m-1, k), \quad (24)$$

where  $q(k, j)$  is the probability of moving from the node  $(m-1, k)$  to  $(m, j)$ . **Note:**  $q(k, j)$  is only non-zero for at most three  $k$ .

- Second, with  $G(m, j)$  at hand, the  $(m+1)$ -period bond price is given by

$$\begin{aligned} P(m+1) &= \sum_{j=-n_m}^{n_m} G(m, j)p(m, j) \\ &= e^{-\alpha_m \Delta} \sum_{j=-n_m}^{n_m} G(m, j)e^{-j(\Delta x)\Delta} \end{aligned} \quad (25)$$

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## Trinomial lattices – 7

- The solution to (25) is readily available in closed form:

$$\alpha_m = \frac{\log \left( \sum_{j=-n_m}^{n_m} G(m, j)e^{-j(\Delta x)\Delta} \right) - \log P(m+1)}{\Delta} \quad (26)$$

- Having found  $\alpha_m$ , we proceed to  $m+1$  (next period) using the same recursions — forward equation (24) followed by (26).
- This completes the construction of the Hull-White trinomial tree for the extended Vasicek model.
- The parameters  $\kappa$  and  $\sigma$  can be calibrated to, e.g., cap prices by minimizing the squared pricing errors

$$S(\kappa, \sigma) = \sum_i \left( V_i^{\text{actual}} - V_i^{\text{model}} \right)^2. \quad (27)$$

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